

Importance Sampling and Statistical Romberg method

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Outline of The Talk

- 1 Introduction
- 2 Robbins-Monro Algorithms
- 3 Central limit theorem for the adaptative procedure
- 4 Numerical results for the Heston model

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- 2 Robbins-Monro Algorithms
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The Model

- $X \in \mathbb{R}^d$ be solution to

$$dX_t = b(X_t)dt + \sum_{j=1}^q \sigma_j(X_t)dW_t^j, \quad X_0 = x \in \mathbb{R}^d$$

where $W = (W^1, \dots, W^q)$ is a q -dimensional Brownian motion.

- Functions $b : \mathbb{R}^d \rightarrow \mathbb{R}^d$ and $\sigma_j : \mathbb{R}^d \rightarrow \mathbb{R}^d$, $1 \leq j \leq q$, satisfy condition

$$(\mathcal{H}_{b,\sigma}) \quad \forall x, y \in \mathbb{R}^d \quad |b(x) - b(y)| + \sum_{j=1}^q |\sigma_j(x) - \sigma_j(y)| \leq C_{b,\sigma}|x - y|,$$

Discretization error

Let X^n be the Euler scheme with time step $\delta = T/n$

$$dX_t^n = b(X_{\eta_n(t)})dt + \sum_{j=1}^q \sigma_j(X_{\eta_n(t)})dW_t^j, \quad \eta_n(t) = [t/\delta]\delta.$$

under condition $(\mathcal{H}_{b,\sigma})$ we have property

$$(\mathcal{P}) \quad \forall p \geq 1, X, X^n \in L^p \quad \text{and} \quad \mathbb{E}\left[\sup_{0 \leq t \leq T} |X_t - X_t^n|^p \right] \leq \frac{K_p(T)}{n^{p/2}}.$$

Discretization error

In the context of possibly degenerate diffusions

- For a given function ψ , we set

$$\varepsilon_n := \mathbb{E}\psi(X_T) - \mathbb{E}\psi(X_T^n)$$

- If ψ , b and σ_j are \mathcal{C}_P^4 then $\varepsilon_n \simeq 1/n$
- However, if ψ is only of class \mathcal{C}^1 , then we have $\varepsilon_n \simeq 1/n^\alpha$ for any $\alpha \in [1/2, 1]$

From now on, we suppose

$$(\mathcal{H}_{\varepsilon_n}^\alpha) \quad n^\alpha \varepsilon_n := n^\alpha \mathbb{E}\psi(X_T^n) - \mathbb{E}\psi(X_T) \rightarrow C_\psi(T, \alpha) \quad \text{for } \alpha \in [1/2, 1].$$

CLT for Monte Carlo method

Theorem

Let $\psi \in \mathcal{C}^1$ s.t. we have

$$(\mathcal{H}_{\varepsilon_n}^\alpha) \quad \lim_{n \rightarrow \infty} n^\alpha \varepsilon_n = C_\psi(T, \alpha)$$

Then,

$$n^\alpha \left(\frac{1}{n^{2\alpha}} \sum_{i=1}^{n^{2\alpha}} \psi(X_T^n, i) - \mathbb{E} \psi(X_T) \right) \Rightarrow \sigma G + C_\psi(T, \alpha),$$

with $\sigma^2 = \text{Var}(\psi(X_T))$.

CLT for Monte Carlo method

Theorem

Let $\psi \in \mathcal{C}^1$ s.t. we have

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with $\sigma^2 = \text{Var}(\psi(X_T))$.

Optimal time complexity

$$C_{MC} = C \times n^{2\alpha+1}$$

Statistical Romberg algorithm

- We construct two Euler schemes X_T^n and $X_T^{\sqrt{n}}$ with time step T/n and T/\sqrt{n} .

- Let

$$E = \mathbb{E}\psi\left(X_T^{\sqrt{n}}\right).$$

- We set

$$Q = \psi\left(X_T^n\right) - \psi\left(X_T^{\sqrt{n}}\right) + E$$

- Note that

$$\mathbb{E}(Q) = \mathbb{E}\psi\left(X_T^n\right) \text{ and } \text{Var}(Q) = O\left(\frac{1}{\sqrt{n}}\right)$$

Statistical Romberg method

The statistical Romberg routine that approximates $\mathbb{E}\psi(X_T)$ using only two empirical means

$$V_n := \frac{1}{N_1} \sum_{i=1}^{N_1} \psi(\hat{X}_{T,i}^{\sqrt{n}}) + \frac{1}{N_2} \sum_{i=1}^{N_2} \psi(X_{T,i}^n) - \psi(X_T^{\sqrt{n}}).$$

Under assumption $(\mathcal{H}_{\varepsilon_n}^\alpha)$, this method is tamed by a central limit theorem with a rate of convergence equal to n^α (Kebaier 2005). More precisely, for $N_1 = n^{2\alpha}$, $N_2 = n^{2\alpha-1/2}$ we have

$$n^\alpha (V_n - \mathbb{E}\psi(X_T)) \rightarrow \mathcal{N}(C_\psi(T, \alpha), \sigma^2), \text{ with}$$

$$\sigma^2 := \text{Var}(\psi(X_T)) + \tilde{\text{Var}}(\nabla\psi(X_T) \cdot U_T),$$

Statistical Romberg method

- The process U is the weak limit process of the error $\sqrt{n}(X^n - X)$ and is solution to

$$dU_t = \dot{b}(X_t)U_t dt + \sum_{j=1}^q \dot{\sigma}_j(X_t)U_t dW_t^j - \frac{1}{\sqrt{2}} \sum_{j,\ell=1}^q \dot{\sigma}_j(X_t)\sigma_\ell(X_t) d\tilde{W}_t^{\ell j},$$

where \tilde{W} is a q^2 -dimensional standard Brownian motion, independent of W , and \dot{b} (respectively $(\dot{\sigma}_j)_{1 \leq j \leq q}$) is the Jacobian matrix of b (respectively $(\sigma_j)_{1 \leq j \leq q}$).

- This result is due to Jacod-Kurtz-Protter (91-98) provided that b and σ are \mathcal{C}^1 .

Importance Sampling

We define the family of \mathbb{P}_θ , as all the equivalent probability measures with respect to \mathbb{P} such that

$$L_t^\theta = \frac{d\mathbb{P}_\theta}{d\mathbb{P}} \Big|_{\mathcal{F}_t} = \exp \left(\theta \cdot W_t - \frac{1}{2} |\theta|^2 t \right).$$

Hence, $B_t^\theta := W_t - \theta t$ is a Brownian motion under \mathbb{P}_θ . This leads to

$$\mathbb{E} \psi(X_T) = \mathbb{E}_\theta \left[\psi(X_T) e^{-\theta \cdot B_T^\theta - \frac{1}{2} |\theta|^2 T} \right].$$

The optimal θ parameter is chosen so that it reduces

$$\text{Var}_\theta \left[\psi(X_T) e^{-\theta \cdot B_T^\theta - \frac{1}{2} |\theta|^2 T} \right]$$

Let us introduce the process X_t^θ solution, under \mathbb{P} , to

$$dX_t^\theta = \left(b(X_t^\theta) + \sum_{j=1}^q \theta_j \sigma_j(X_t^\theta) \right) dt + \sum_{j=1}^q \sigma_j(X_t^\theta) dW_t^j,$$

$(B_t^\theta, X_t)_{t \geq 0}$ under \mathbb{P}_θ has the same law as $(W_t, X_t^\theta)_{t \geq 0}$ under \mathbb{P} we get

$$\mathbb{E}\psi(X_T) = \mathbb{E}g(\theta, X_T^\theta, W_T), \text{ with } g(\theta, x, y) = \psi(x)e^{-\theta \cdot y - \frac{1}{2}|\theta|^2 T}.$$

We also introduce the Euler continuous approximation $X^{n,\theta}$ of the process X^θ solution, under \mathbb{P} , to

$$dX_t^{n,\theta} = \left(b(X_{\eta_n(t)}^{n,\theta}) + \sum_{j=1}^q \theta_j \sigma_j(X_{\eta_n(t)}^{n,\theta}) \right) dt + \sum_{j=1}^q \sigma_j(X_{\eta_n(t)}^\theta) dW_t^j,$$

Our target now is to approximate $\mathbb{E}\psi(X_T) = \mathbb{E}g(\theta, X_T^\theta, W_T)$ by

$$\frac{1}{N_1} \sum_{i=1}^{N_1} g(\theta, \hat{X}_{T,i}^{\sqrt{n},\theta}, \hat{W}_{T,i}) + \frac{1}{N_2} \sum_{i=1}^{N_2} g(\theta, X_{T,i}^{\sqrt{n},\theta}, W_{T,i}) - g(\theta, X_{T,i}^{\sqrt{n},\theta}, W_{T,i}).$$

According to Kebaier (2005), we have a CLT with limit variance

$$\text{Var} \left(g(\theta, X_T^\theta, W_T) \right) + \tilde{\text{Var}} \left(\nabla_x g(\theta, X_T^\theta, W_T) \cdot U_T^\theta \right)$$

where U^θ is the weak limit process of the error $\sqrt{n}(X^{n,\theta} - X^\theta)$, solution to

$$dU_t^\theta = \left(b(X_t^\theta) + \sum_{j=1}^q \theta_j \dot{\sigma}_j(X_t^\theta) \right) U_t^\theta dt + \sum_{j=1}^q \dot{\sigma}_j(X_t^\theta) U_t^\theta dW_t^j - \frac{1}{\sqrt{2}} \sum_{j,\ell=1}^q \dot{\sigma}_j(X_t^\theta) \sigma_\ell(X_t^\theta) d\tilde{W}_t^{\ell j}.$$

it boils down to choose $\theta^* = \operatorname{argmin}_{\theta \in \mathbb{R}^q} v(\theta)$

$$v(\theta) := \tilde{\mathbb{E}} \left(\left[\psi(X_T^\theta)^2 + (\nabla \psi(X_T^\theta) \cdot U_T^\theta)^2 \right] e^{-2\theta \cdot W_T - |\theta|^2 T} \right)$$

Note that $v(\theta)$ is not explicit, we introduce $\theta_n^* := \operatorname{argmin}_{\theta \in \mathbb{R}^q} v_n(\theta)$

$$v_n(\theta) := \tilde{\mathbb{E}} \left(\left[\psi(X_T^{n,\theta})^2 + (\nabla \psi(X_T^{n,\theta}) \cdot U_T^{n,\theta})^2 \right] e^{-2\theta \cdot W_T - |\theta|^2 T} \right)$$

with $U^{n,\theta}$ is the Euler discretization scheme of U^θ , solution to

$$\begin{aligned} dU_t^{n,\theta} &= \left(b(X_{\eta_n(t)}^{n,\theta}) + \sum_{j=1}^q \theta_j \dot{\sigma}_j(X_{\eta_n(t)}^{n,\theta}) \right) U_{\eta_n(t)}^{n,\theta} dt \\ &+ \sum_{j=1}^q \dot{\sigma}_j(X_{\eta_n(t)}^{n,\theta}) U_{\eta_n(t)}^{n,\theta} dW_t^j - \frac{1}{\sqrt{2}} \sum_{j,\ell=1}^q \dot{\sigma}_j(X_{\eta_n(t)}^{n,\theta}) \sigma_\ell(X_{\eta_n(t)}^{n,\theta}) d\tilde{W}_t^{\ell j}. \end{aligned}$$

Theorem

Suppose σ and b are in \mathcal{C}^2 with bounded first and second derivatives. Then for any $\theta \in \mathbb{R}$ the following property holds

$$(\tilde{\mathcal{P}}) \forall p \geq 1, U^\theta, U^{n,\theta} \in L^p \text{ and } \tilde{\mathbb{E}} \left[\sup_{0 \leq t \leq T} |U_t^\theta - U_t^{n,\theta}|^p \right] \leq \frac{K_p(T)}{n^{p/2}}.$$

In particular, for $\theta = 0$ the above property holds for the processes U and U^n .

existence and uniqueness of θ^*

Theorem

Suppose σ and b are in \mathcal{C}^2 with bounded first and second derivatives and let ψ in \mathcal{C}^1 such that $\mathbb{P}(\psi(X_T) \neq 0) > 0$.

- If there exists $a > 1$ such that $\mathbb{E} [\psi^{2a}(X_T)]$ and $\mathbb{E} [|\nabla\psi(X_T)|^{2a}]$ are finite,

Then the function $\theta \mapsto v(\theta)$ is \mathcal{C}^2 and strictly convex with $\nabla v(\theta) = \tilde{\mathbb{E}}H(\theta, X_T, U_T, W_T)$ where

$$H(\theta, X_T, U_T, W_T) := (\theta T - W_T) [\psi(X_T)^2 + (\nabla\psi(X_T) \cdot U_T)^2] e^{-\theta \cdot W_T + \frac{1}{2}|\theta|^2 T}.$$

Moreover, there exists a unique $\theta^* \in \mathbb{R}^q$ such that $\min_{\theta \in \mathbb{R}^q} v(\theta) = v(\theta^*)$.

Proof

First of all, note the process (B, X, U) under $\tilde{\mathbb{P}}_\theta$ has the same law as (W, X^θ, U^θ) under $\tilde{\mathbb{P}}$. So, using a change of probability, we get

$$v(\theta) := \tilde{\mathbb{E}} \left(\left[\psi(X_T)^2 + (\nabla \psi(X_T) \cdot U_T)^2 \right] e^{-\theta \cdot W_T + \frac{1}{2} |\theta|^2 T} \right).$$

It follows that

- The map $\theta \mapsto \left[\psi(X_T)^2 + (\nabla \psi(X_T) \cdot U_T)^2 \right] e^{-\theta \cdot W_T + \frac{1}{2} |\theta|^2 T}$ is \mathcal{C}^1
- $\nabla v(\theta) = H(\theta, X_T, U_T, W_T)$

For $c > 0$ we have,

$$\sup_{|\theta| \leq c} |H(\theta, X_T, U_T, W_T)| \leq (cT + |W_T|) \left[\psi(X_T)^2 + (\nabla \psi(X_T) \cdot U_T)^2 \right] e^{c|W_T| + \frac{1}{2} c^2 T}.$$

Proof

- Using Holder's inequality, $\tilde{\mathbb{E}} \sup_{|\theta| \leq c} |H(\theta, X_T, U_T, W_T)|$ is bounded by

$$e^{\frac{1}{2}c^2T} \left(\|\psi^2(X_T)\|_a \|e^{c|W_T|}(cT + |W_T|)\|_{\frac{a}{a-1}} + \|\nabla\psi(X_T)\|_a \|U_T\|_{\frac{2a}{a-1}} \|e^{c|W_T|}(cT + |W_T|)\|_{\frac{2a}{a-1}} \right).$$

- Using property $(\tilde{\mathcal{P}})$ and $\mathbb{E}\psi^{2a}(X_T)$ and $\mathbb{E}|\nabla\psi(X_T)|^{2a}$ are finite we conclude the boundedness of $\tilde{\mathbb{E}} \sup_{|\theta| \leq c} |H(\theta, X_T, U_T, W_T)|$.

- In the same way, we prove that v is of class \mathcal{C}^2 in \mathbb{R}^q

$$\begin{aligned} \text{Hess}(v(\theta)) &= \tilde{\mathbb{E}} \left[((\theta T - W_T)(\theta T - W_T)^* + T I_q) \right. \\ &\quad \left. \times (\psi^2(X_T) + (\nabla\psi(X_T) \cdot U_T)^2) e^{-\theta \cdot W_T + \frac{1}{2}|\theta|^2 T} \right]. \end{aligned}$$

Since $\mathbb{P}(\psi(X_T) \neq 0) > 0$, we get for all $u \in \mathbb{R}^q \setminus \{0\}$

$$\begin{aligned} u^* \text{Hess}(v(\theta)) u &= \tilde{\mathbb{E}} \left[T|u|^2 + (u \cdot (\theta T - W_T))^2 (\psi^2(X_T) \right. \\ &\quad \left. + (\nabla\psi(X_T) \cdot U_T)^2) e^{-\theta \cdot W_T + \frac{1}{2}|\theta|^2 T} \right] > 0. \end{aligned}$$

- Now it will be sufficient to prove that $\lim_{|\theta| \rightarrow \infty} v(\theta) = +\infty$

$$v(\theta) = \tilde{\mathbb{E}} \left[(\psi(X_T)^2 + (\nabla\psi(X_T) \cdot U_T)^2) e^{-\theta \cdot W_T + \frac{1}{2}|\theta|^2 T} \right].$$

$$\begin{aligned} +\infty &= \tilde{\mathbb{E}} \left[\liminf_{|\theta| \rightarrow \infty} (\psi(X_T)^2 + (\nabla\psi(X_T) \cdot U_T)^2) e^{-\theta \cdot W_T + \frac{1}{2}|\theta|^2 T} \right] \\ &\leq \liminf_{|\theta| \rightarrow +\infty} \tilde{\mathbb{E}} \left[(\psi(X_T)^2 + (\nabla\psi(X_T) \cdot U_T)^2) e^{-\theta \cdot W_T + \frac{1}{2}|\theta|^2 T} \right]. \end{aligned}$$

The same results can be obtained for the Euler scheme X^n .

Theorem

Suppose σ and b are in \mathcal{C}^2 with bounded first and second derivatives. Let ψ be \mathcal{C}^1 such that $\mathbb{P}(\psi(X_T^n) \neq 0) > 0$.

- If there exists $a > 1$ such that $\mathbb{E}[\psi^{2a}(X_T^n)]$ and $\mathbb{E}[|\nabla\psi(X_T^n)|^{2a}]$ are finite

Then the function $\theta \mapsto v_n(\theta)$ is \mathcal{C}^2 and strictly convex with

$$\nabla v_n(\theta) = \tilde{\mathbb{E}}H(\theta, X_T^n, U_T^n, W_T).$$

- Moreover, there exists a unique $\theta_n^* \in \mathbb{R}^q$ such that $\min_{\theta \in \mathbb{R}^q} v_n(\theta) = v_n(\theta_n^*)$.

Further, we prove the convergence of θ_n^* towards θ^* as n tends to infinity.

Theorem

Suppose σ and b are in \mathcal{C}^2 with bounded first and second derivatives. Let ψ be \mathcal{C}^1 such that $\mathbb{P}(\psi(X_T) \neq 0) > 0$ and for all $n \in \mathbb{N}$, $\mathbb{P}(\psi(X_T^n) \neq 0) > 0$.

- If there exists $a > 1$ such that $\mathbb{E}[\psi^{2a}(X_T)]$, $\sup_{n \in \mathbb{N}} \mathbb{E}[\psi^{2a}(X_T^n)]$, $\mathbb{E}[|\nabla \psi(X_T)|^{2a}]$ and $\sup_{n \in \mathbb{N}} \mathbb{E}[|\nabla \psi(X_T^n)|^{2a}]$ are finite.

Then,

$$\theta_n^* \longrightarrow \theta^*, \quad \text{as } n \rightarrow \infty.$$

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- The aim now is to construct for fixed n some sequences $(\theta_i^n)_{i \in \mathbb{N}}$ such that $\lim_{i \rightarrow \infty} \theta_i^n = \theta_n^* \arg \min_{\theta \in \mathbb{R}} v_n(\theta) =$ almost surely.
- Indeed, using the Robbins-Monro algorithm, we construct recursively the sequence of random variables $(\theta_i^n)_{i \in \mathbb{N}}$ in \mathbb{R}^q given by

$$\theta_{i+1}^n = \theta_i^n - \gamma_{i+1} H(\theta_i^n, X_{T,i+1}^n, U_{T,i+1}^n, W_{T,i+1}), \quad i \geq 0, \quad \theta_0^n \in \mathbb{R}^q,$$

$(\gamma_i)_{i \geq 1}$ is a decreasing sequence of positive real numbers satisfying

$$\sum_{i=1}^{\infty} \gamma_i = \infty \quad \text{and} \quad \sum_{i=1}^{\infty} \gamma_i^2 < \infty$$

- To obtain the a.s. convergence of θ_i^n to θ_n^* , we need to check
 - $\forall \theta \neq \theta_n^*, \langle \nabla v_n(\theta), \theta - \theta_n^* \rangle > 0,$
 - (NEC) $\tilde{\mathbb{E}} [|H(\theta, X_T^n, U_T^n, W_T)|^2] \leq C(1+|\theta|^2),$ for all $\theta \in \mathbb{R}^q.$

Unfortunately, this condition is not satisfied in our context.

Constrained stochastic algorithm

Let $(\mathcal{K}_i)_{i \in \mathbb{N}}$ denote an increasing sequence of compact sets satisfying $\bigcup_{i=0}^{\infty} \mathcal{K}_i = \mathbb{R}^d$ and $\mathcal{K}_i \subsetneq \overset{\circ}{\mathcal{K}}_{i+1}, \forall i \in \mathbb{N}$. For $\theta_0^n \in \mathcal{K}_0$, $\alpha_0^n = 0$ and a gain sequence $(\gamma_i)_{i \in \mathbb{N}}$ satisfying (??), we define the sequence $(\theta_i^n, \alpha_i^n)_{i \in \mathbb{N}}$ recursively by

$$\left\{ \begin{array}{l} \text{if } \theta_i^n - \gamma_{i+1} H(\theta_i^n, X_{T,i+1}^n, U_{T,i+1}^n, W_{T,i+1}) \in \mathcal{K}_{\alpha_i^n}, \text{ then} \\ \theta_{i+1}^n = \theta_i^n - \gamma_{i+1} H(\theta_i^n, X_{T,i+1}^n, U_{T,i+1}^n, W_{T,i+1}), \text{ and } \alpha_{i+1}^n = \alpha_i^n \\ \text{else } \theta_{i+1}^n = \theta_i^n \text{ and } \alpha_{i+1}^n = \alpha_i^n + 1, \end{array} \right.$$

Constrained stochastic algorithm

Theorem

Suppose σ and b are \mathcal{C}^2 with bounded first and second derivatives and ψ is \mathcal{C}^1 . Assume that for all $n \in \mathbb{N}$, $\mathbb{P}(\psi(X_T^n) \neq 0) > 0$.

- there exists $a > 1$ s.t. $\mathbb{E} [\psi^{4a}(X_T^n)]$ and $\mathbb{E} [|\nabla \psi(X_T^n)|^{4a}] < \infty$

Then the sequence $(\theta_i^n)_{i \geq 0}$ satisfies

- 1 For all $n \in \mathbb{N}$, we have $\theta_i^n \xrightarrow{i \rightarrow \infty} \theta_n^*$, a.s.
- 2 Reversely, for all $i \in \mathbb{N}$, we have $\theta_i^n \xrightarrow{n \rightarrow \infty} \theta_i$, a.s.,

$$\left\{ \begin{array}{l} \text{if } \theta_i - \gamma_{i+1} H(\theta_i, X_{T,i+1}, U_{T,i+1}, W_{T,i+1}) \in \mathcal{K}_{\alpha_i}, \text{ then} \\ \theta_{i+1} = \theta_i - \gamma_{i+1} H(\theta_i, X_{T,i+1}, U_{T,i+1}, W_{T,i+1}), \text{ and } \alpha_{i+1} = \alpha_i \\ \text{else } \theta_{i+1} = \theta_0 \text{ and } \alpha_{i+1} = \alpha_i, \end{array} \right.$$

The following corollary follows immediately

Corollary

Under above assumptions the constrained algorithm given satisfies

$$\lim_{i,n \rightarrow \infty} \theta_i^n = \lim_{i \rightarrow \infty} \left(\lim_{n \rightarrow \infty} \theta_i^n \right) = \lim_{n \rightarrow \infty} \left(\lim_{i \rightarrow \infty} \theta_i^n \right) = \theta^*, \quad \tilde{\mathbb{P}}\text{-a.s.},$$

where $\theta^* = \operatorname{argmin}_{\theta \in \mathbb{R}^q} v(\theta)$

$$v(\theta) := \tilde{\mathbb{E}} \left(\left[\psi(X_T^\theta)^2 + (\nabla \psi(X_T^\theta) \cdot U_T^\theta)^2 \right] e^{-2\theta \cdot W_T - |\theta|^2 T} \right)$$

Unconstrained stochastic algorithm

- We use the idea proposed by Lemaire and Pagès (2009), a new algorithm that satisfies (NEC). In our context we have

$$\nabla v_n(\theta) = \tilde{\mathbb{E}} \left((\theta T - W_T) [\psi(X_T^n)^2 + (\nabla \psi(X_T^n) \cdot U_T^n)^2] e^{-\theta \cdot W_T + \frac{1}{2} |\theta|^2 T} \right).$$

- To do so, we apply Girsanov theorem, with shift parameter $-\theta$.

$$B_t^{(-\theta)} := W_t + \theta t \quad \text{and} \quad L_t^{(-\theta)} := \frac{d\mathbb{P}_{(-\theta)}}{d\mathbb{P}} \Big|_{\mathcal{F}_t} = e^{-\theta \cdot W_t - \frac{1}{2} |\theta|^2 t}$$

$$\begin{aligned} \nabla v_n(\theta) &= \tilde{\mathbb{E}}_{(-\theta)} \left[(2\theta T - B_T^{(-\theta)}) [\psi(X_T^n)^2 + (\nabla \psi(X_T^n) \cdot U_T^n)^2] e^{|\theta|^2 T} \right] \\ &= \tilde{\mathbb{E}} \left[(2\theta T - W_T) [\psi(X_T^{n,(-\theta)})^2 + (\nabla \psi(X_T^{n,(-\theta)}) \cdot U_T^{n,(-\theta)})^2] e^{|\theta|^2 T} \right], \end{aligned}$$

since $(B^{(-\theta)}, X^n, U^n, \tilde{\mathbb{P}}_{(-\theta)}) \stackrel{\text{law}}{=} (W, X^{n,(-\theta)}, U^{n,(-\theta)}, \tilde{\mathbb{P}})$

- We need in this context to strengthen our assumptions on ψ and suppose that $\partial_\alpha \psi$ are with polynomial growth for $|\alpha| \leq 1$
- we introduce for a given $\eta > 0$, a new function

$$\begin{aligned} \tilde{H}_\eta(\theta, X_T^{n,(-\theta)}, U_T^{n,(-\theta)}, W_T) &= e^{-\eta|\theta|^2 T} (2\theta T - W_T) \\ &\quad \times \left[\psi(X_T^{n,(-\theta)})^2 + (\nabla \psi(X_T^{n,(-\theta)})) \cdot U_T^{n,(-\theta)} \right]^2. \end{aligned}$$

Then, the algorithm is given by

$$\theta_{i+1}^n = \theta_i^n - \gamma_{i+1} H_\eta(\theta_i^n, X_{T,i+1}^{n,(-\theta_i^n)}, U_{T,i+1}^{n,(-\theta_i^n)}, W_{T,i+1}), \quad \theta_0 \in \mathbb{R}. \quad (1)$$

This algorithm would behave like a classical Robbins-Monro one and does not suffer from the violation of (NEC).

Theorem

Suppose σ and b are \mathcal{C}^2 with bounded first and second derivatives. Let ψ in \mathcal{C}^1 such that for and for all $n \in \mathbb{N}$, $\mathbb{P}(\psi(X_T^n) \neq 0) > 0$. In addition, assume that for $\lambda > 0$ we have

$$|\nabla\psi(x)| \leq C_\psi(1 + |x|^\lambda) \quad \text{for all } x \in \mathbb{R}^d \text{ and } C_\psi > 0.$$

Then, the sequence $(\theta_i^n)_{i \geq 0}$ given by routine (1), satisfies

$$\forall n \in \mathbb{N}, \quad \theta_i^n \xrightarrow{i \rightarrow \infty} \theta_n^*, \quad \text{a.s.}$$

$$\theta_n^* := \operatorname{argmin}_{\theta \in \mathbb{R}^q} \tilde{\mathbb{E}} \left(\left[\psi(X_T^{n,\theta})^2 + (\nabla\psi(X_T^{n,\theta}) \cdot U_T^{n,\theta})^2 \right] e^{-2\theta \cdot W_T - |\theta|^2 T} \right)$$

Proof

- We have to check first that $\forall \theta \neq \theta_n^*$

$$\langle h_n(\theta), \theta - \theta_n^* \rangle > 0, \text{ where } h_n(\theta) = \tilde{\mathbb{E}} H_\eta(\theta, X_T^{n,(-\theta)}, U_T^{n,(-\theta)}, W_T).$$

This is immediate since $h_n(\theta) = K_\eta(\theta) \nabla v_n(\theta)$ with $K_\eta > 0$.

- It remains to prove

$$\sup_{\theta \in \mathbb{R}^q} \tilde{\mathbb{E}} \left[|H_\eta(\theta, X_T^{n,(-\theta)}, U_T^{n,(-\theta)}, W_T)|^2 \right] < \infty,$$

By Cauchy-Schwartz inequality we obtain for $\lambda_1 = 4\lambda \vee 2(\lambda + 1)$,

$$\begin{aligned} \tilde{\mathbb{E}} \left[|H_\eta(\theta, X_T^{n,(-\theta)}, U_T^{n,(-\theta)}, W_T)|^2 \right] &\leq e^{-2\eta|\theta|^2 T} \left\| |2\theta T - W_T|^2 \right\|_2 \\ &\quad \times \left(\left\| \psi(X_T^{n,(-\theta)})^2 \right\|_2 + \left\| (\nabla \psi(X_T^{n,(-\theta)}) \cdot U_T^{n,(-\theta)})^2 \right\|_2 \right). \end{aligned}$$

$$\begin{aligned} \tilde{\mathbb{E}} \left[|H_\eta(\theta, X_T^{n,(-\theta)}, U_T^{n,(-\theta)}, W_T)|^2 \right] \\ \leq C e^{-2\eta|\theta|^2 T} (1 + |\theta|^2) \left(1 + \left\| |X_T^{n,(-\theta)}|^{\lambda_1} \right\|_2 + \left\| |U_T^{n,(-\theta)}|^4 \right\|_2 \right). \end{aligned}$$

- Using properties (\mathcal{P}) and $(\tilde{\mathcal{P}})$, we get

$$\begin{aligned} \tilde{\mathbb{E}} \left[|H_\eta(\theta, X_T^{n,(-\theta)}, U_T^{n,(-\theta)}, W_T)|^2 \right] &\leq C e^{-2\eta|\theta|^2 T} (1 + |\theta|^2) \\ &\times \left(1 + \left\| |X_T^{n,(-\theta)} - X_T^n|^{\lambda_1} \right\|_2 + \left\| |U_T^{n,(-\theta)} - U_T^n|^4 \right\|_2 \right) \end{aligned}$$

- Using Gronwall inequality, we obtain that

$$\tilde{\mathbb{E}} \left| X_T^{n,(-\theta)} - X_T^n \right|^{2\lambda_1} \leq C |\theta|^{2\lambda_1} \sum_{j=1}^q \tilde{\mathbb{E}} \left| \int_0^T |\sigma_j(X_s^{n,(-\theta)})| ds \right|^{2\lambda_1},$$

$$\tilde{\mathbb{E}} \left| U_T^{n,(-\theta)} - U_T^n \right|^8 \leq C |\theta|^8 \tilde{\mathbb{E}} \left| \int_0^T |U_s^{n,(-\theta)}| ds \right|^8.$$

- As $(B^{(-\theta)}, X^n, U^n, \tilde{\mathbb{P}}_{(-\theta)}) \stackrel{law}{=} (W, X^{n,(-\theta)}, U^{n,(-\theta)}, \tilde{\mathbb{P}})$,

$$\tilde{\mathbb{E}} \left| \int_0^T |\sigma_j(X_s^{n,(-\theta)})| ds \right|^{2\lambda_1} = \tilde{\mathbb{E}} \left(\left| \int_0^T |\sigma_j(X_s^n)| ds \right|^{2\lambda_1} e^{-\theta \cdot W_T - \frac{1}{2} |\theta|^2 T} \right)$$

In the same way

$$\tilde{\mathbb{E}} \left| \int_0^T |U_s^{n,(-\theta)}| ds \right|^8 = \tilde{\mathbb{E}} \left(\left| \int_0^T |U_s^n| ds \right|^8 e^{-\theta \cdot W_T - \frac{1}{2} |\theta|^2 T} \right).$$

Now using Holder's inequality, with $\frac{1}{r} + \frac{1}{r'} = 1$, the linear growth of $(\sigma_j)_{1 \leq j \leq q}$, properties (\mathcal{P}) and $(\tilde{\mathcal{P}})$ we obtain

$$\begin{aligned} \tilde{\mathbb{E}} \left[|H_\eta(\theta, X_T^{n,(-\theta)}, U_T^{n,(-\theta)}, W_T)|^2 \right] \\ \leq C e^{-2\eta|\theta|^2 T} (1 + |\theta|^2) \left(1 + (|\theta|^{\lambda_1} + |\theta|^4) e^{\frac{r-1}{4} |\theta|^2 T} \right). \end{aligned}$$

We complete the proof by choosing $r \in]1, 1 + 8\eta[$. □

Theorem

Suppose b and σ are \mathcal{C}^2 with bounded first and second derivatives and s.t. all the derivatives of order 2 are Lipschitz continuous. Let $\nabla\psi$ in \mathcal{C}^1 s.t. for all $n \in \mathbb{N}$, $\mathbb{P}(\psi(X_T^n) \neq 0) > 0$.

- Assume also that $\partial_\alpha\psi$, for $|\alpha| \leq 2$, are with polynomial growth

Then, for all $\forall i \in \mathbb{N}$, $p \geq 1$, there exists $C_i > 0$ such that

$$\forall n \in \mathbb{N}^*, \quad \tilde{\mathbb{E}}|\theta_{i+1}^n - \theta_{i+1}|^{2p} \leq \frac{C_i}{n^p}.$$

Moreover,

$$\forall i \in \mathbb{N}, \quad \theta_i^n \xrightarrow[n \rightarrow \infty]{} \theta_i, \quad \text{a.s.}$$

where the sequence $(\theta_i)_{i \geq 0}$ is introduced in the above corollary.

Proof

We give the proof in the case of dimension one and proceed by induction on $i \in \mathbb{N}$. For all $p \geq 1$ relation (1) yields

$$\begin{aligned} \tilde{\mathbb{E}}(\theta_{i+1}^n - \theta_{i+1})^{2p} &\leq C \tilde{\mathbb{E}}(\theta_i^n - \theta_i)^{2p} \\ &\quad + C \gamma_{i+1}^{2p} \tilde{\mathbb{E}} \left(H_\eta(\theta_i^n, X_{T,i+1}^{n,(-\theta_i^n)}, U_{T,i+1}^{n,(-\theta_i^n)}, W_{T,i+1}) \right. \\ &\quad \left. - H_\eta(\theta_i, X_{T,i+1}^{(-\theta_i)}, U_{T,i+1}^{(-\theta_i)}, W_{T,i+1}) \right)^{2p}. \end{aligned}$$

Using the induction assumption we only need to control the second term bounded by $C \times (\tilde{\mathbb{E}}H_1^{2p} + \tilde{\mathbb{E}}H_2^{2p})$

$$\begin{aligned} H_1 &:= e^{-\eta|\theta_i^n|^2 T} (2\theta_i^n T - W_{T,i+1}) \\ &\quad \times \left[\psi(X_{T,i+1}^{n,(-\theta_i^n)})^2 + \left(\psi'(X_{T,i+1}^{n,(-\theta_i^n)}) U_{T,i+1}^{n,(-\theta_i^n)} \right)^2 \right. \\ &\quad \left. - \psi(X_{T,i+1}^{(-\theta_i)})^2 - \left(\psi'(X_{T,i+1}^{(-\theta_i)}) U_{T,i+1}^{(-\theta_i)} \right)^2 \right] \end{aligned}$$

$$H_2 := e^{-\eta|\theta_i^n|^2 T} (2\theta_i^n T - W_{T,i+1}) \left[\psi(X_{T,i+1}^{(-\theta_i^n)})^2 + \left(\psi'(X_{T,i+1}^{(-\theta_i^n)}) U_{T,i+1}^{(-\theta_i^n)} \right)^2 \right] \\ - e^{-\eta|\theta_i|^2 T} (2\theta_i T - W_{T,i+1}) \left[\psi(X_{T,i+1}^{(-\theta_i)})^2 + \left(\psi'(X_{T,i+1}^{(-\theta_i)}) U_{T,i+1}^{(-\theta_i)} \right)^2 \right].$$

• We give the proof for H_2 . Here, we need first to introduce, for all $u \in \mathbb{R}$, the couple of u -sensitivity processes $(Y_t^{(-u)}, Z_t^{(-u)})_{t \in [0, T]}$

given by $Y_t^{(-u)} := \frac{\partial X_t^{(-u)}}{\partial u}$ and $Z_t^{(-u)} := \frac{\partial U_t^{(-u)}}{\partial u}$, $t \in [0, T]$

Therefore, we write $\tilde{\mathbb{E}} H_2^{2p} = \tilde{\mathbb{E}} B(\theta_i^n, \theta_i)$

$$B(\theta, \theta') = \tilde{\mathbb{E}} \left[\int_{(\theta', \theta)} \frac{\partial}{\partial u} \left\{ e^{-\eta|u|^2 T} (2uT - W_T) \right. \right. \\ \left. \left. \times \left[\psi(X_T^{(-u)})^2 + \left(\psi'(X_T^{(-u)}) U_T^{(-u)} \right)^2 \right] \right\} du \right]^{2p}.$$

- Now, $B(\theta, \theta') \leq C \sum_{i=1}^4 B_i(\theta, \theta')$. These four terms are of the same type, so we only treat one of them let's say B_3

$$B_3(\theta, \theta') \leq C |\theta - \theta'|^{2p-1} \times \int_{(\theta', \theta)} \tilde{\mathbb{E}} \left[(2uT - W_T)^{2p} e^{-2p\eta|u|^2 T} (U_T^{(-u)})^{2p} \psi'(X_T^{(-u)})^{4p} (Z_T^{(-u)})^{2p} \right] du$$

- Note that the same probability change leading to cancel the u -term in the drift part of $X^{(-u)}$ operates in the same way for the other processes $U^{(-u)}$, $Y^{(-u)}$ and $Z^{(-u)}$. So for all $u \in \mathbb{R}$, we get $(B^{(-u)}, X, U, Y, Z, \tilde{\mathbb{P}}_{(-u)}) \stackrel{law}{=} (W, X^{(-u)}, U^{(-u)}, Y^{(-u)}, Z^{(-u)}, \tilde{\mathbb{P}})$

$$B_3(\theta, \theta') \leq C |\theta - \theta'|^{2p-1} \int_{(\theta', \theta)} \tilde{\mathbb{E}} \left[(uT - W_T)^{2p} e^{-2p\eta|u|^2 T} \times (U_T)^{2p} \psi'(X_T)^{4p} (Z_T)^{2p} e^{-uW_T - \frac{1}{2}|u|^2 T} \right] du.$$

$$B_3(\theta, \theta') \leq C|\theta - \theta'|^{2p},$$

- This is immediate, since X and U satisfy properties (\mathcal{P}) and $(\tilde{\mathcal{P}})$ and Z is a diffusion process with enough smooth coefficients satisfying likewise the same type of properties.
- So that, we obtain for all $p > 1$

$$B(\theta, \theta') \leq C|\theta - \theta'|^{2p}.$$

Now, since $\tilde{\mathbb{E}}H_2^{2p} = \tilde{\mathbb{E}}B(\theta_i^n, \theta_i)$, it follows for all $p > 1$

$$\tilde{\mathbb{E}}H_2^{2p} \leq C\mathbb{E}|\theta_i^n - \theta_i|^{2p} \leq \frac{C_i}{n^p}.$$

Corollary

Under above assumptions if $\mathbb{P}(\psi(X_T) \neq 0) > 0$, then the unconstrained algorithm satisfies

$$\lim_{i,n \rightarrow \infty} \theta_i^n = \lim_{i \rightarrow \infty} \left(\lim_{n \rightarrow \infty} \theta_i^n \right) = \lim_{n \rightarrow \infty} \left(\lim_{i \rightarrow \infty} \theta_i^n \right) = \theta^*, \quad \tilde{\mathbb{P}}\text{-a.s.},$$

where $\theta^* = \underset{\theta \in \mathbb{R}^q}{\operatorname{argmin}} v(\theta)$

$$v(\theta) := \tilde{\mathbb{E}} \left(\left[\psi(X_T^\theta)^2 + (\nabla \psi(X_T^\theta) \cdot U_T^\theta)^2 \right] e^{-2\theta \cdot W_T - |\theta|^2 T} \right)$$

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Adaptive Statistical Romberg method

- The adaptive importance sampling algorithm for the statistical Romberg method approximates our initial quantity of interest

$$\mathbb{E}\psi(X_T) = \mathbb{E} \left[\psi(X_T^\theta) e^{-\theta \cdot W_T - \frac{1}{2}|\theta|^2 T} \right] \text{ by}$$

$$\begin{aligned} & \frac{1}{N_1} \sum_{i=1}^{N_1} g(\hat{\theta}_i^m, \hat{X}_{T,i+1}^m, \hat{W}_{T,i+1}) \\ & + \frac{1}{N_2} \sum_{i=1}^{N_2} \left(g(\theta_i^n, X_{T,i+1}^n, W_{T,i+1}) - g(\theta_i^n, X_{T,i+1}^m, W_{T,i+1}) \right), \end{aligned}$$

where for all $x \in \mathbb{R}^d$ and $y \in \mathbb{R}^q$, $g(\theta, x, y) = \psi(x) e^{-\theta \cdot y - \frac{1}{2}|\theta|^2 T}$.

- Here the paths generated by W and \hat{W} are of course independent.

Lindeberg Feller CLT

Theorem

Suppose that $(\Omega, \mathbb{F}, \mathbb{P})$ is a probability space and that for each n , we have a filtration $\mathbb{F}_n = (\mathcal{F}_k^n)_{k \geq 0}$, a sequence $k_n \rightarrow \infty$ as $n \rightarrow \infty$ and a real square integrable vector martingale $M^n = (M_k^n)_{k \geq 0}$ which is adapted to \mathbb{F}_n such that

- There exists a deterministic symmetric positive semi-definite matrix Γ , such that

$$\langle M \rangle_{k_n}^n = \sum_{k=1}^{k_n} \mathbb{E} [|M_k^n - M_{k-1}^n|^2 | \mathcal{F}_{k-1}^n] \xrightarrow[n \rightarrow \infty]{\mathbb{P}} \Gamma.$$

- There exists a real number $a > 1$, such that

$$\sum_{k=1}^{k_n} \mathbb{E} [|M_k^n - M_{k-1}^n|^{2a} | \mathcal{F}_{k-1}^n] \xrightarrow[n \rightarrow \infty]{\mathbb{P}} 0.$$

Then

$$M_{k_n}^n \xrightarrow{\mathcal{L}} \mathcal{N}(0, \Gamma) \quad \text{as } n \rightarrow \infty.$$

Toeplitz Lemma

Lemma

Let $(a_i)_{1 \leq i \leq k_n}$ a sequence of real positive numbers, where $k_n \uparrow \infty$ as n tends to infinity, and $(x_i^n)_{i \geq 1, n \geq 1}$ a double indexed sequence such that

$$(i) \quad \lim_{n \rightarrow \infty} \sum_{1 \leq i \leq k_n} a_i = \infty$$

$$(ii) \quad \lim_{i, n \rightarrow \infty} x_i^n = \lim_{i \rightarrow \infty} \left(\lim_{n \rightarrow \infty} x_i^n \right) = \lim_{n \rightarrow \infty} \left(\lim_{i \rightarrow \infty} x_i^n \right) = x < \infty$$

Then

$$\lim_{n \rightarrow +\infty} \frac{\sum_{i=1}^{k_n} a_i x_i^n}{\sum_{i=1}^{k_n} a_i} = x.$$

The adaptive Monte Carlo method

Theorem

- Let $(\theta_i^n)_{i \geq 0}$, $n \in \mathbb{N}$ and $(\theta_i)_{i \geq 0}$ satisfying (\mathcal{H}_θ) $\lim_{i, n \rightarrow \infty} \theta_i^n = \lim_{i \rightarrow \infty} (\lim_{n \rightarrow \infty} \theta_i^n) = \lim_{n \rightarrow \infty} (\lim_{i \rightarrow \infty} \theta_i^n) = \theta^*$, $\tilde{\mathbb{P}}$ -a.s.,
- Assume that b and σ satisfy $(\mathcal{H}_{b, \sigma})$ and the function ψ is a real valued function satisfying assumption $(\mathcal{H}_{\varepsilon_n})$, with $\alpha \in [1/2, 1]$ and $C_\psi \in \mathbb{R}$, s.t. $|\psi(x) - \psi(y)| \leq C(1 + |x|^p + |y|^p)|x - y|$, then the following convergence holds

$$n^\alpha \left(\frac{1}{n^{2\alpha}} \sum_{i=1}^{n^{2\alpha}} g(\theta_i^n, X_{T, i+1}^{n, \theta_i^n}, W_{T, i+1}) - \mathbb{E}\psi(X_T) \right) \xrightarrow{\mathcal{L}} \mathcal{N}(C_\psi, \sigma^2).$$

$$\sigma^2 := \mathbb{E} \left(\psi(X_T)^2 e^{-\theta^* \cdot W_T - \frac{1}{2} |\theta^*|^2 T} \right) - [\mathbb{E}\psi(X_T)]^2$$

The adaptive statistical Romberg method

Theorem

- Let $(\theta_i^n)_{i \geq 0}$, $n \in \mathbb{N}$ and $(\theta_i)_{i \geq 0}$ satisfying (\mathcal{H}_θ) .
- Assume that b and σ are \mathcal{C}^1 satisfying $(\mathcal{H}_{b,\sigma})$ and ψ is \mathcal{C}^1 , satisfying $(\mathcal{H}_{\varepsilon_n})$, with constants $\alpha \in (1/2, 1]$ and $C_\psi \in \mathbb{R}$, s.t.

$$|\psi(x) - \psi(y)| \leq C(1 + |x|^p + |y|^p)|x - y|, \quad \text{for some } C, p > 0.$$

If we choose $N_1 = n^{2\alpha}$, $N_2 = n^{2\alpha - \beta}$ then

$$n^\alpha (V_n - \mathbb{E}\psi(X_T)) \xrightarrow{\mathcal{L}} \mathcal{N}(C_\psi, \sigma^2 + \tilde{\sigma}^2) \quad \text{as } n \rightarrow \infty,$$

where $\sigma^2 = \mathbb{E} \left[\psi(X_T)^2 e^{-\theta^* \cdot W_T - \frac{1}{2} |\theta^*|^2 T} \right] - [\mathbb{E}\psi(X_T)]^2$ and

$$\tilde{\sigma}^2 := \tilde{\mathbb{E}} \left[[\nabla \psi(X_T) \cdot U_T]^2 e^{-\theta^* \cdot W_T - \frac{1}{2} |\theta^*|^2 T} \right].$$

Proof

- We prove only the convergence of the second empirical mean in the Statistical Romberg method. To do so, we introduce the martingale arrays $(M_k^n)_{k \geq 1}$. For $\beta = 1/2$

$$M_k^n := \frac{1}{n^{\alpha-\beta}} \sum_{i=1}^k \left(g(\theta_i^n, X_{T,i+1}^{n,\theta_i^n}, W_{T,i+1}) - g(\theta_i^n, X_{T,i+1}^{n,\beta}, W_{T,i+1}) - \mathbb{E}[\psi(X_T^n) - \psi(X_T^{n,\beta})] \right),$$

- The quadratic variation of M evaluated at $n^{2\alpha-\beta}$ is given by

$$\langle M \rangle_{n^{2\alpha-\beta}}^n = \frac{1}{n^{2\alpha-\beta}} \sum_{i=1}^{n^{2\alpha-\beta}} n^\beta \xi_n(\theta_i^n) - \left(n^{\frac{\beta}{2}} [\mathbb{E}\psi(X_T^n) - \mathbb{E}\psi(X_T^{n,\beta})] \right)^2,$$

where $\forall \theta \in \mathbb{R}^q$, $\xi_n(\theta) := \mathbb{E} \left([\psi(X_T^n) - \psi(X_T^{n,\beta})]^2 e^{-\theta \cdot W_T + \frac{1}{2} |\theta|^2 T} \right)$.

- We focus now on the asymptotic behavior of $n^\beta \xi_n(\theta)$. Applying Taylor expansion theorem twice we get for all $\theta \in \mathbb{R}^q$

$$\begin{aligned} n^{\frac{\beta}{2}} [\psi(X_T^n) - \psi(X_T^{n^\beta})] e^{-\frac{1}{2}\theta \cdot W_T + \frac{1}{4}|\theta|^2 T} \\ = n^{\frac{\beta}{2}} \nabla \psi(X_T) \cdot [X_T^n - X_T^{n^\beta}] e^{-\frac{1}{2}\theta \cdot W_T + \frac{1}{4}|\theta|^2 T} + R_n, \end{aligned}$$

$$R_n := n^{\frac{\beta}{2}} (X_T^n - X_T) \varepsilon(X_T, X_T^n - X_T) - n^{\frac{\beta}{2}} (X_T^{n^\beta} - X_T) \varepsilon(X_T, X_T^{n^\beta} - X_T)$$

with $\varepsilon(X_T, X_T^n - X_T) \xrightarrow{\mathbb{P}\text{-a.s.}} 0$ and $\varepsilon(X_T, X_T^{n^\beta} - X_T) \xrightarrow{\mathbb{P}\text{-a.s.}} 0$

- Further, as b and σ are \mathcal{C}^1 functions then we have the tightness of $n^{\frac{\beta}{2}} (X_T^n - X_T)$ and $n^{\frac{\beta}{2}} (X_T^{n^\beta} - X_T)$ and we deduce that $R_n \rightarrow 0$.

- It follows

$$n^{\frac{\beta}{2}} [\psi(X_T^n) - \psi(X_T^{n^\beta})] e^{-\frac{1}{2}\theta \cdot W_T + \frac{1}{4}|\theta|^2 T} \xrightarrow{\text{stably}} \nabla \psi(X_T) \cdot U_T e^{-\frac{1}{2}\theta \cdot W_T + \frac{1}{4}|\theta|^2 T}.$$

- Otherwise, $\forall \theta \in \mathbb{R}^q$ and $a' > 1$ we have thanks to the assumption on ψ together with property (\mathcal{P}) , we obtain

$$\sup_n \mathbb{E} \left| n^{\frac{\beta}{2}} [\psi(X_T^n) - \psi(X_T^{n^\beta})] e^{-\frac{1}{2}\theta \cdot W_T + \frac{1}{4}|\theta|^2 T} \right|^{2a'} < \infty$$

So, $\lim_{n \rightarrow \infty} n^\beta \xi_n(\theta) = \tilde{\mathbb{E}} \left([\nabla \psi(X_T) \cdot U_T]^2 e^{-\theta \cdot W_T + \frac{1}{2}|\theta|^2 T} \right) := \xi(\theta)$

- Using property (\mathcal{P}) with assumption on ψ , we check the equicontinuity of the family functions $(n^\beta \xi_n)_{n \geq 1}$.
- So under assumption (\mathcal{H}_θ) , we get

$$\lim_{i, n \rightarrow \infty} n^\beta \xi_n(\theta_i^n) = \xi(\theta^*) \quad \tilde{\mathbb{P}}\text{-a.s.}$$

Then, Toeplitz Lemma yields $\lim_{n \rightarrow \infty} \langle M \rangle_{n^{2\alpha - \beta}}^n = \xi(\theta^*)$, $\tilde{\mathbb{P}}\text{-a.s.}$

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The popular stochastic volatility model in finance is the Heston model solution to

$$\begin{cases} dS_t = rS_t dt + \sqrt{V_t} S_t dW_t^1 \\ dV_t = \kappa(\bar{v} - V_t) dt + \sigma \sqrt{V_t} \rho dW_t^1 + \sigma \sqrt{V_t} \sqrt{1 - \rho^2} dW_t^2, \end{cases}$$

where W^1 and W^2 are two independent Brownian motions.

Parameters $\kappa, \sigma, \bar{v}, r > 0$ and $|\rho| \leq 1$.

- Our aim is to use the importance sampling method in order to reduce the variance when computing the price is

$$e^{-rT} \mathbb{E} \psi(S_T) = e^{-rT} \mathbb{E} \left[g(\theta, S_T^\theta) \right] = e^{-rT} \mathbb{E} \left[\psi(S_T^\theta) e^{-\theta \cdot W_T - \frac{1}{2} |\theta|^2 T} \right],$$

To approximate S_T^θ , we consider the step T/n and we discretize the stochastic process using the Euler scheme. For $i \in [0, n-1]$ and $\theta = (\theta_1, \theta_2) \in \mathbb{R}^2$,

$$\begin{cases} S_{t_{i+1}}^{n,\theta} = S_{t_i}^{n,\theta} \left(1 + (r + \theta_1 \sqrt{V_{t_i}^{n,\theta}}) \frac{T}{n} + \sqrt{V_{t_i}^{n,\theta}} \frac{T}{n} Z_{1,i+1} \right), \\ V_{t_{i+1}}^{n,\theta} = \left| V_{t_i}^{n,\theta} + \left(\kappa(\bar{v} - V_{t_i}^{n,\theta}) \right. \right. \\ \left. \left. + \sigma \sqrt{V_{t_i}^{n,\theta}} (\rho\theta_1 + \sqrt{1-\rho^2}\theta_2) \right) \frac{T}{n} + \sigma \sqrt{V_{t_i}^{n,\theta}} \frac{T}{n} Z_{2,i+1} \right|, \end{cases}$$

Hence, the price is firstly approximated by

$$e^{-rT} \mathbb{E} \left[g(\theta, S_T^{n,\theta}) \right] = e^{-rT} \mathbb{E} \left[\psi(S_T^{n,\theta}) e^{-\theta \cdot W_T - \frac{1}{2} |\theta|^2 T} \right], \quad \theta \in \mathbb{R}^2.$$

- The optimal θ for a Monte Carlo method

$$\theta_n^* = \arg \min_{\theta \in \mathbb{R}^2} \mathbb{E} \left[\psi^2(S_T^{n,\theta}) e^{-2\theta \cdot W_T - |\theta|^2 T} \right].$$

- The optimal θ for the second one is

$$\tilde{\theta}_n^* = \arg \min_{\theta \in \mathbb{R}^2} \mathbb{E} \left[\left(\psi^2(S_T^{n,\theta}) + (\nabla \psi(S_T^{n,\theta}) \cdot U_T^{n,\theta})^2 \right) e^{-2\theta \cdot W_T - |\theta|^2 T} \right],$$

- Here, we have also the choice of the algorithm approximating both θ_n^* and $\tilde{\theta}_n^*$ by the constrained algorithm or by the unconstrained algorithm
- We fix $S_0 = 100$, $V_0 = 0.01$, $K = 100$, the free interest rate $r = \log(1.1)$, $\sigma = 0.2$, $k = 2$, $\bar{v} = 0.01$, $\rho = 0.5$ and maturity time $T = 1$.

	Constrained algorithm	Unconstrained algorithm
θ_n^*	(0.7906, 0.0516)	(0.7904, 0.0532)
$\tilde{\theta}_n^*$	(0.7884, 0.0587)	(0.7898, 0.0576)

Table: Estimation of θ_n^* and $\tilde{\theta}_n^*$

- 1 First, we choose $\gamma_i = \gamma_0/i^\alpha$, for $\alpha \in (\frac{1}{2}, 1)$ and $\gamma_0 > 0$.
- 2 Then, we compute $\bar{\theta}_{i+1}^n := \frac{1}{i+1} \sum_{k=0}^i \tilde{\theta}_k^n$.

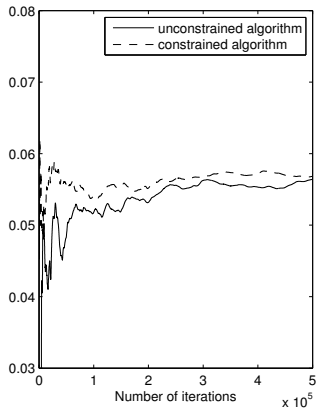
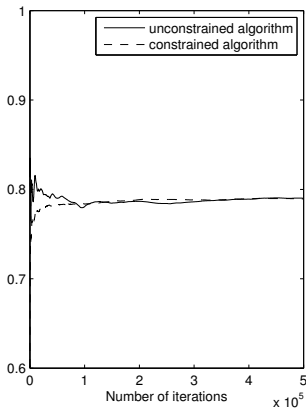


Figure: Values obtained with $n = 100$, $\gamma_0 = 0.01$ and $\alpha = 0.75$.

Our aim now, is to compare

- MC+IS method: European call option price approximation with $N = n^2$

$$\frac{e^{-rT}}{N} \sum_{i=1}^N g(\theta_M^n, S_{T,i+1}^{n, \theta_M^n})$$

- SR+IS method: European call option price approximation method with $N_1 = n^2$ and $N_2 = n^{\frac{3}{2}}$

$$\begin{aligned} & \frac{e^{-rT}}{N_1} \sum_{i=1}^{N_1} g(\tilde{\theta}_M^n, \hat{S}_{T,i+1}^{\sqrt{n}, \tilde{\theta}_M^n}) \\ & + \frac{e^{-rT}}{N_2} \sum_{i=1}^{N_2} \left(g(\tilde{\theta}_M^n, S_{T,i+1}^{n, \tilde{\theta}_M^n}) - g(\tilde{\theta}_M^n, S_{T,i+1}^{\sqrt{n}, \tilde{\theta}_M^n}) \right). \end{aligned}$$

Method	n	Price	CI length	time
MC+IS	400	9.641444	0.060094	10.38
	900	9.661192	0.029409	91.5
	1600	9.656892	0.016538	512.29
SR+IS	600	9.659409	0.057454	3.36
	1600	9.660062	0.019933	26.79
	3600	9.65673	0.008584	194.6

Table: Call Price for the Heston model

Method	n	Price	CI length	time
MC+IS	400	0.863968	0.00721	9.39
	900	0.863291	0.003151	91.58
	1600	0.863766	0.001774	515.31
SR+IS	600	0.867441	0,007249	3.27
	1600	0.864213	0.002541	27.02
	3600	0.862589	0.001095	202.2

Table: Delta call price for the Heston model

The first method (MC+IS) is already implemented and available in the free online version of Premia platform (<https://www.rocq.inria.fr/mathfi/Premia/index.html>) and our method (SR+IS) is now added in the latest premium version.

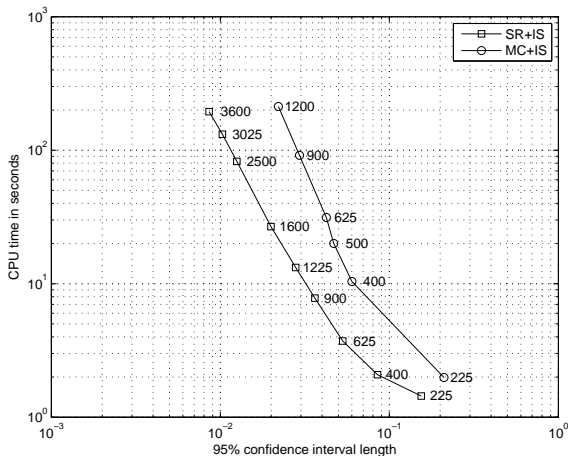


Figure: CPU time versus the 95%-confidence interval length

Thank you!

Multilevel Monte Carlo and the Euler scheme

- We use $L + 1$ Euler schemes with time steps $\frac{T}{m^\ell}$ for $\ell \in \{0, 1, \dots, L\}$ such that $m^L = n$.
- We can write

$$\mathbb{E}(f(X_T^n)) = \mathbb{E}\left(f(X_T^{m^0})\right) + \sum_{\ell=1}^L \mathbb{E}\left(f(X_T^{m^\ell}) - f(X_T^{m^{\ell-1}})\right).$$

- The Multilevel method consists on estimating independently each of the expectations above.

$$Q_n = \frac{1}{N_0} \sum_{k=1}^{N_0} f(X_{T,k}^{m^0}) + \sum_{\ell=1}^L \frac{1}{N_\ell} \sum_{k=1}^{N_\ell} \left(f(X_{T,k}^{m^\ell}) - f(X_{T,k}^{m^{\ell-1}})\right).$$

we have

$$\text{Var}(Q_n) = O\left(\sum_{\ell=1}^L N_\ell^{-1} m^{-\ell}\right).$$